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Addressing continuous missing outcomes in pairwise and network meta-analysis

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ERC Grant IMMA 260559



Why missing data matter

- Missing outcome data are common in RCT's
- In mental health, the dropout rate may exceed 50%.
- Loss in power and precision
- Any analysis must make an untestable assumption about missing data
 - wrong assumptions  biased estimates
- Some popular techniques for missing data get biased standard errors.
 - resulting in wrong p-values and confidence intervals

Reasons why data are missing

- Missing At Random (MAR)

The probability that data are missing

- may depend on the values of the observed data
- does not depend on the values of the missing data (conditional on the values of the observed data)

- Missing Not At Random (MNAR) or Informatively Missing (IM)

The probability that data are missing

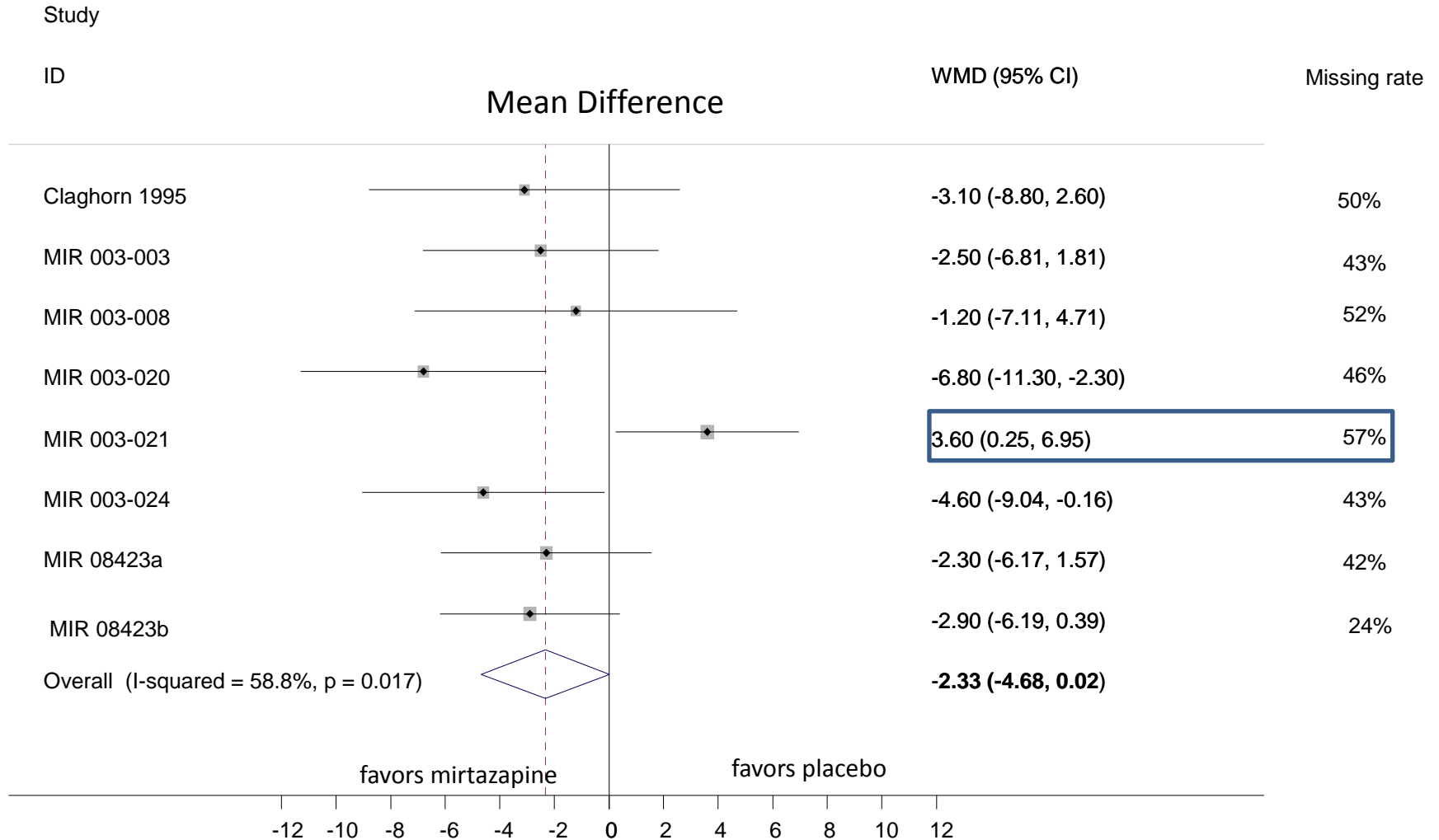
- may depend on unobserved data
- may depend on the actual missing outcome

- The widely used approach is to assume that missing data are MAR conditional on the treatment received. This justifies a complete case (CC) analysis (unbiased results with lower precision)
- **We consider a single quantitative(continuous) outcome measured at one time point without considering any covariates**

Mirtazapine meta-analysis: mean change in HAM21 scores, standard deviations, and numbers of completers and non-completers for the mirtazapine and placebo arms.

Study	Placebo				Mirtazapine			
	x_{iC}^{obs}	sd_{iC}^{obs}	n_{iC}	m_{iC}	x_{iT}^{obs}	sd_{iT}^{obs}	n_{iT}	m_{iT}
Claghorn 1995	-11.4	10.2	19	26	-14.5	8.8	26	19
MIR 003-003	-11.5	8.3	24	21	-14	7.3	27	18
MIR 003-008	-11.4	8	17	13	-12.6	8	12	18
MIR 003-020	-6.2	6.5	24	19	-13	9	23	21
MIR 003-021	-17.4	5.3	21	29	-13.8	5.9	22	28
MIR 003-024	-11.1	9.9	27	23	-15.7	6.7	30	20
MIR 84023a	-11.9	8.6	33	24	-14.2	7.6	35	25
MIR 84023b	-11.8	8.3	48	18	-14.7	8.4	51	13

Random effect meta-analysis of mean change in HAM-D21 score. Mirtazapine vs placebo. Complete case analysis

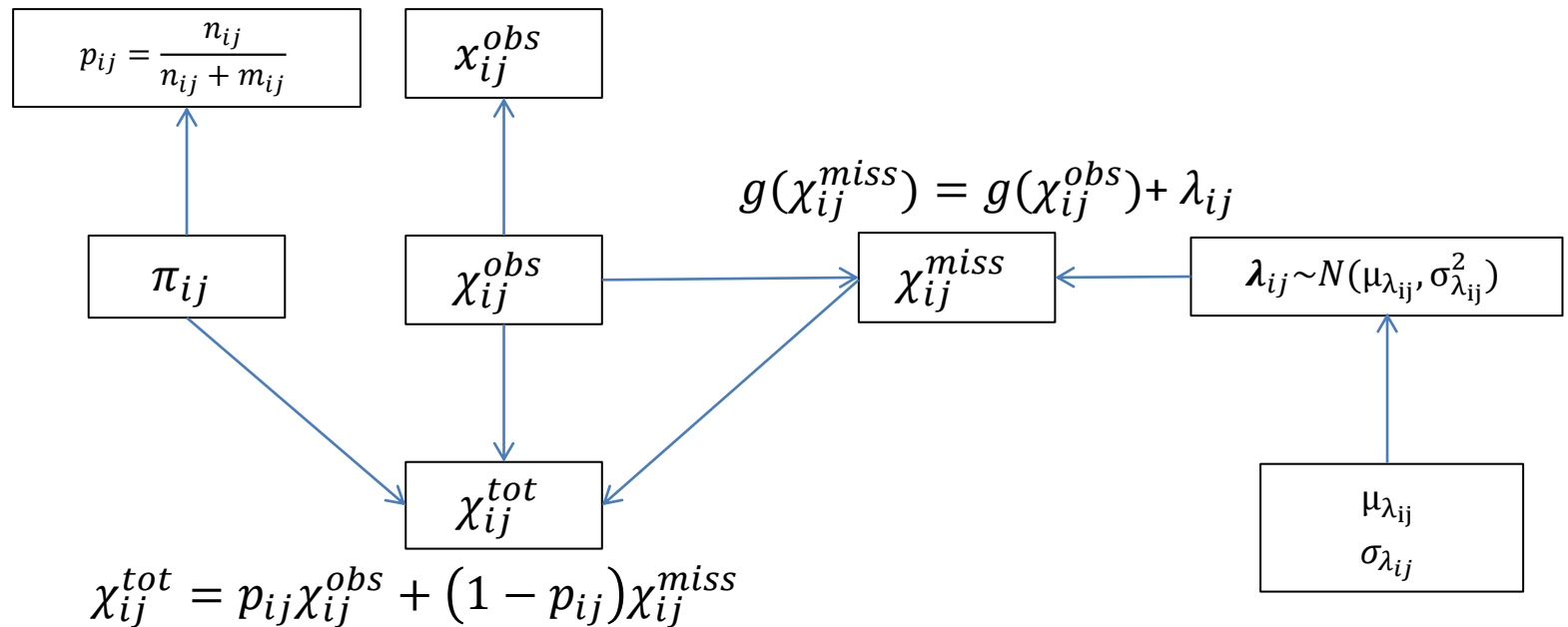


Our proposal for continuous missing outcomes

Two-stage method

- We develop a two-stage method that models departures from MAR
- 1st stage: estimate effect sizes β_i and its variance $var(\beta_i)$ in each study i .
 - increase uncertainty to effect sizes to account for the fact that missing data are imputed rather than observed.
 - downweight studies with large missing rates
- 2nd stage: Pool adjusted effect sizes

Model for arm j of study i pattern mixture model



$$x_{ij}^{tot} = p_{ij}x_{ij}^{obs} + (1 - p_{ij})g^{-1}((\lambda_{ij} + g(x_{ij}^{obs})))$$

Parameters for Informative Missingness λ

Difference of Means (IMDoM) or Ratio of Means (IMRoM)

- IM is quantified as a contrast between the unobserved mean value in the missing data and the mean value in the observed data

$$\lambda_{ij} = g(\chi_{ij}^{miss}) - g(\chi_{ij}^{obs})$$

1. If g is the identity function (IMDoM)

$$\lambda_{ij} = \chi_{ij}^{miss} - \chi_{ij}^{obs}$$

2. If g is the \ln function (IMRoM)

$$\lambda_{ij} = \ln(\chi_{ij}^{miss}) - \ln(\chi_{ij}^{obs})$$

Assumptions about the missingness parameter

- Missing at random (MAR) $\lambda_{ij} = 0$
- Free λ_{ij} : $\lambda_{ij} \sim N(\mu_{\lambda_{ij}}, \sigma_{\lambda_{ij}}^2)$
- Study specific λ : $\lambda_i \sim N(\mu_{\lambda_i}, \sigma_{\lambda_i}^2)$

- Correlated λ 's:

$$\begin{pmatrix} \lambda_{iC} \\ \lambda_{iT} \end{pmatrix} \sim N \left(\begin{pmatrix} \mu_{\lambda_{iC}} \\ \mu_{\lambda_{iT}} \end{pmatrix}, \begin{pmatrix} \sigma_{\lambda_{iC}}^2 & \rho_{\lambda} \sigma_{\lambda_{iC}} \sigma_{\lambda_{iT}} \\ \rho_{\lambda} \sigma_{\lambda_{iC}} \sigma_{\lambda_{iT}} & \sigma_{\lambda_{iT}}^2 \end{pmatrix} \right)$$

Adjusted effect sizes

$$\beta_i = f(x_{iT}^{tot}) - f(x_{iC}^{tot})$$

1. $f(u_i) = u_i$: Mean Difference (MD)
2. $f(u_i) = \ln u_i$: Logarithm of the ratio of means (ln RoM)
3. $f(u_i) = u_i/S_i$: Standardized Mean Difference (SMD) where S_i is the pooled standard deviation

Estimating $E(\beta_i)$ and $var(\beta_i)$

Taylor series approximation

$$\beta_i = f(x_{iT}^{tot}) - f(x_{iC}^{tot})$$

$$x_{ij}^{tot} = \pi_{ij}x_{ij}^{obs} + (1 - \pi_{ij})g^{-1}\left((\lambda_{ij} + g(x_{ij}^{obs}))\right)$$

- $E(\beta_i)$ is straightforward calculated, e.g. if f and g are identity functions (MD and IMDOM)

$$E(\beta_i) = (1 - p_{iT})\mu_{\lambda_{iT}} + x_{iT}^{obs} - (1 - p_{iC})\mu_{\lambda_{iC}} - x_{iC}^{obs}$$

$$var(\beta_i) \approx \sum_{j=C,T} \left[\frac{p_{ij}(1 - p_{ij})}{n_{ij} + m_{ij}} (\mu_{\lambda_{ij}}^2 + \sigma_{\lambda_{ij}}^2) + \frac{s_{ij}^2}{n_{ij}} + \sigma_{\lambda_{ij}}^2 (1 - p_{ij})^2 \right] - 2\rho_{\lambda_i} \sigma_{\lambda_{iC}} \sigma_{\lambda_{iT}} (1 - p_{iT})(1 - p_{iC})$$

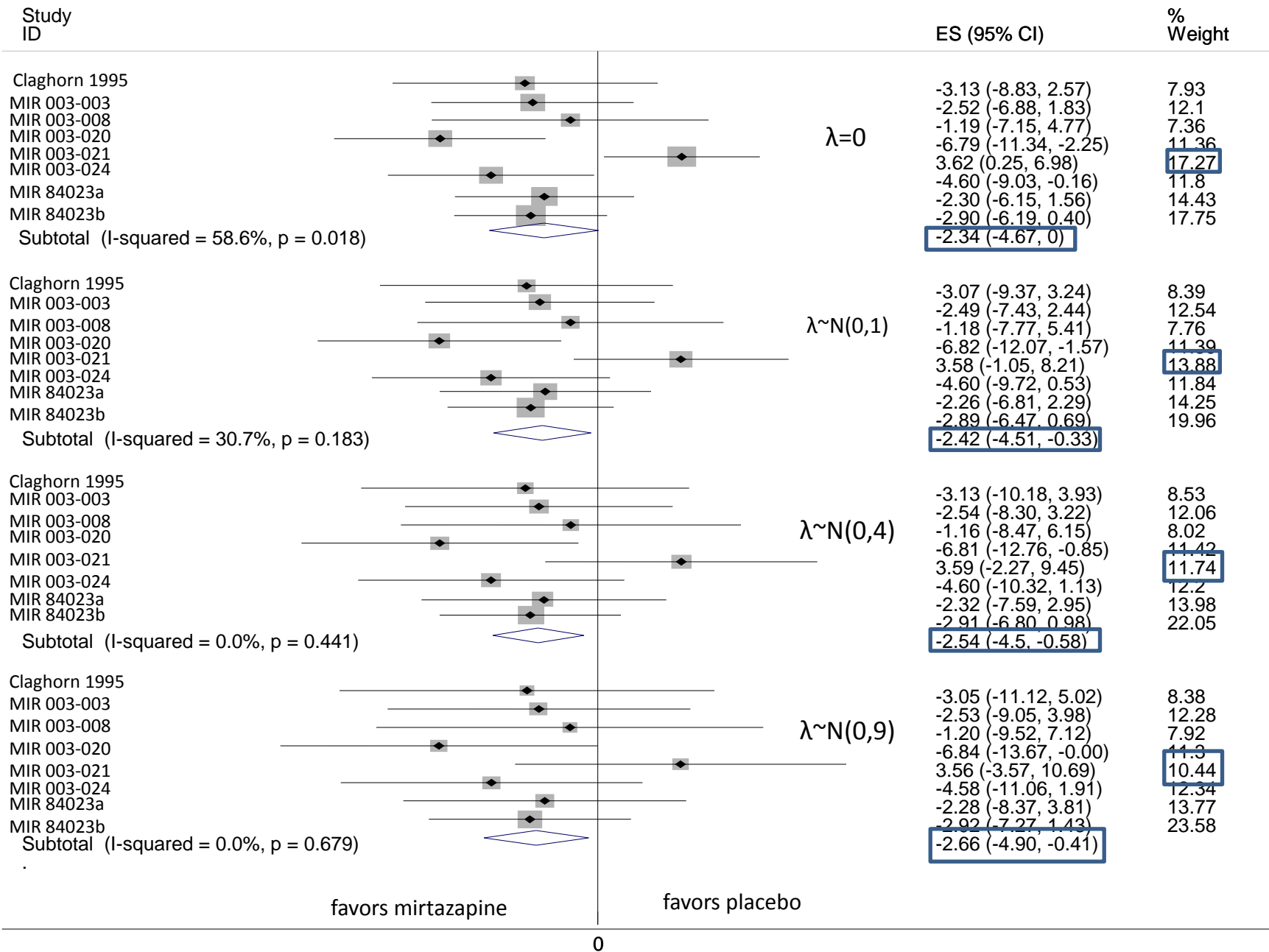
Estimating $E(\beta_i)$ and $var(\beta_i)$

Monte Carlo Method

- Repeat B times, $b = 1, \dots, B$
- draw $(\chi_{ij}^{obs})^b \sim N\left(x_{ij}^{obs}, \frac{s_{ij}^2}{n_{ij}}\right)$
- draw $\lambda_{ij}^b \sim N\left(\mu_{\lambda_{ij}}, \sigma_{\lambda_{ij}}^2\right)$
- set $(\chi_{ij}^{miss})^b = \lambda_{ij}^b + (\chi_{ij}^{obs})^b$ for IMDoM or
 $(\chi_{ij}^{miss})^b = e^{\lambda_{ij}^b} (\chi_{ij}^{obs})^b$ for IMRoM
- draw $\pi_{ij}^b \sim N\left(p_{ij}, \frac{p_{ij}(1-p_{ij})}{n_{ij}}\right) T(0,1)$
- compute $(\chi_{ij}^{tot})^b = \pi_{ij}^b (\chi_{ij}^{obs})^b + (1 - \pi_{ij}^b) (\chi_{ij}^{miss})^b$
- Compute $E(\beta_i)$ and $var(\beta_i)$

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Conclusions/Future work

- We developed a model that can
 - model departures from MAR (IMDoF,IMRoM) for MD, SMD and RoM
 - downweight studies with lots of missing data
- we need priors for IMDoM/IMRoM or conduct a sensitivity analysis
- there is a trade-off between within-study and between-study variability
- method extended to network meta-analysis allowing for multi-arm studies
- metamiss command in STATA (in collaboration with Anna Chaimani)

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